

# Dynamic robust Orlicz premia and Haezendonck-Goovaerts risk measures\*

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## Abstract

In this paper we extend to a dynamic setting the robust Orlicz premia and Haezendonck-Goovaerts risk measures introduced in Bellini, Laeven and Rosazza Gianin [4]. We extensively analyze the properties of the resulting dynamic risk measures. Furthermore, we characterize dynamic Orlicz premia that are time-consistent, and establish some relations between the time-consistency properties of dynamic robust Orlicz premia and the corresponding dynamic robust Haezendonck-Goovaerts risk measures.

**Keywords:** Orlicz premia, norms and spaces; Haezendonck-Goovaerts risk measures; Time-consistency; Robustness; Ambiguity averse preferences.

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# 1 Introduction

In the context of risk pricing and capital requirements, modern static theories of financial risk measurement are provided by monetary, convex and entropy convex measures of risk (Föllmer and Schied [24, 25], Frittelli and Rosazza Gianin [27], Ruszczynski and Shapiro [41], and Laeven and Stajje [34]).

Over the past two decades not only the study of static but also of dynamic theories of risk measurement has developed into a flourishing and mathematically refined area of research. The dynamic counterparts of the static theories of monetary measures of risk have been developed in Artzner et al. [1], Riedel [38], Frittelli and Rosazza Gianin [28], Detlefsen and Scandolo [20], Cheridito, Delbaen and Kupper [10], Delbaen [17], Föllmer and Penner [23], Klöppel and Schweizer [32], Cheridito and Kupper [13], among many others. We refer to Föllmer and Schied [25], Chapter 11, for an overview and many references.

A main problem in dynamic risk measurement is the consistency over time of the evaluation as well as of the resulting decisions. The notion of *recursiveness*, or Bellman's dynamic programming principle, has played a central role in the early development of the literature on the theory and application of dynamic measurement of risk; see e.g., Duffie and Epstein [21], Chen and Epstein [9], Epstein and Schneider [22] and Ruszczynski and Shapiro [42]. Recursiveness is intimately related to (strong) time-consistency (even equivalent, under linear utility).

A large literature analyzes and characterizes time-consistency for the canonical theories of risk measurement. There are several approaches to characterizing/generating time-consistency properties in the literature, including the following: approaches based on mixture representations and law invariance (Weber [44], Kupper and Schachermayer [33], Delbaen, Bellini, Bignozzi and Ziegel [19]); based on dual representations, requiring  $m$ -stability (or rectangularity) of the set of generalized scenarios, or imposing the cocycle property on the penalty function (Delbaen [17], Föllmer and Penner [23] and Bion-Nadal [5, 6]); based on a decomposition property of acceptance sets (Cheridito, Delbaen and Kupper [10]); and based on a recursive construction via generators (Cheridito and Kupper [14]). In continuous-time, in a Brownian or Brownian-Poissonian filtration, another approach consists of characterizing time-consistency via a representation of the penalty function; see Delbaen, Peng and Rosazza Gianin [18], Tang and Wei [43], and Laeven and Stajje [35].

Recently, Bellini, Laeven and Rosazza Gianin [4] introduced *robust return risk measures*, mainly for two purposes. First, to reveal and formalize the difference between risk measurement in terms of *monetary* values and in terms of *returns*. Second, to take into account in this setting ambiguity with respect to the probabilistic model  $P$ , by means of ambiguity averse preferences, specifically invoking multiple priors (Gilboa and Schmeidler [29]), variational preferences (Maccheroni, Marinacci and Rustichini [36]), or homothetic preferences (Chateauneuf and Faro [8], Cerreia-Vioglio, Maccheroni, Marinacci and Montrucchio [7] and Laeven and Stajje [34]).

In particular, in Bellini, Laeven and Rosazza Gianin [4] we provided an axio-

matic foundation of a canonical subclass of return risk measures, that of *Orlicz premia*, by exploiting a one-to-one correspondence between Orlicz premia and measures of (utility-based) shortfall risk. Furthermore, we defined, axiomatized and studied robustified versions of Orlicz premia and of their optimized translation-invariant extensions (Rockafellar and Uryasev [39] and Rockafellar, Uryasev and Zabarankin [40]), known as *Haezendonck-Goovaerts risk measures*; see Haezendonck and Goovaerts [31], Delbaen [16], Goovaerts, Kaas, Dhaene and Tang [30] and Bellini and Rosazza Gianin [2, 3] for the classical (non-robust) definitions. We explicated that Orlicz premia can be interpreted to assess the stochastic nature of returns—they are return risk measures—, in contrast to the common use of monetary risk measures to assess the stochastic nature of a position’s monetary value. The class of return risk measures encompasses interesting subclasses of risk measures, such as  $p$ -norms, that are not included in the class of monetary measures of risk.

In this paper we extend to a dynamic setting the static robust return risk measures introduced in Bellini, Laeven and Rosazza Gianin [4]. We extensively analyze the properties of the resulting dynamic robust return risk measures. Furthermore, we provide characterization results of their time-consistency. We show in particular that the only time-consistent dynamic Orlicz premia are conditional  $p$ -norms. We also show that time-consistency of dynamic robust Orlicz premia and of the associated Haezendonck-Goovaerts risk measures are intimately related.

The remainder of this paper is organized as follows: In Section 2 we introduce our notation and setting and recall some preliminaries. In Section 3 we introduce dynamic Orlicz premia and analyze their properties. In Section 4 we consider their robust extension. In Section 5 we introduce dynamic robust Haezendonck-Goovaerts risk measures. In Section 6 we analyze time-consistency properties.

## 2 Preliminaries and basic definitions

Throughout the paper, we work on a nonatomic probability space  $(\Omega, \mathcal{F}, P)$ . All equalities and inequalities between random variables are meant to hold  $P$ -a.s. without further specification. We denote by  $L^\infty(P)$ ,  $L_+^\infty(\mathcal{F}, P)$ , and  $L_{++}^\infty(P)$  the sets of  $P$ -a.s. bounded,  $P$ -a.s. bounded non-negative, and  $P$ -a.s. bounded strictly positive random variables, respectively. We assume that positive realizations of random variables represent losses. A risk measure  $\rho: L^\infty(P) \rightarrow \mathbb{R}$  is said to be:

- monotone, if  $X \leq Y \Rightarrow \rho(X) \leq \rho(Y)$
- translation invariant, if  $\rho(X + h) = \rho(X) + h, \forall h \in \mathbb{R}, \forall X \in L^\infty$
- monetary, if it is monotone, translation invariant and satisfies  $\rho(0) = 0$
- convex, if  $\rho(\alpha X + (1 - \alpha)Y) \leq \alpha\rho(X) + (1 - \alpha)\rho(Y), \forall \alpha \in [0, 1], \forall X, Y \in L^\infty$

- positively homogeneous, if  $\rho(\lambda X) = \lambda\rho(X)$ ,  $\forall \lambda \geq 0, \forall X \in L^\infty$
- subadditive, if  $\rho(X + Y) \leq \rho(X) + \rho(Y)$ ,  $\forall X, Y \in L^\infty$
- coherent, if it is monotone, translation invariant, positively homogeneous and subadditive
- law-invariant, if  $X \stackrel{d}{=} Y \Rightarrow \rho(X) = \rho(Y)$ .

A risk measure  $\rho$  has the Fatou property if

$$X_n \xrightarrow{P} X, \|X_n\|_\infty \leq k \Rightarrow \rho(X) \leq \liminf_{n \rightarrow +\infty} \rho(X_n),$$

while it has the stronger Lebesgue property if

$$X_n \xrightarrow{P} X, \|X_n\|_\infty \leq k \Rightarrow \rho(X) = \lim_{n \rightarrow +\infty} \rho(X_n).$$

With our sign conventions, the Lebesgue property is equivalent to continuity from above, i.e.,

$$X_n \downarrow X \Rightarrow \rho(X_n) \rightarrow \rho(X),$$

while the Fatou property is equivalent to continuity from below, that is,

$$X_n \uparrow X \Rightarrow \rho(X_n) \rightarrow \rho(X).$$

We let  $\Phi: [0, +\infty) \rightarrow [0, +\infty)$  be strictly increasing and convex, with  $\Phi(0) = 0$ ,  $\Phi(1) = 1$ , and  $\Phi(+\infty) = +\infty$ . Such a  $\Phi$  is referred to as a Young function.

**Definition 1.** Let  $\Phi$  be a Young function. For a random loss  $X \in L_+^\infty(\Omega, \mathcal{F}, P)$ , the Orlicz premium is defined by

$$H^\Phi(X) := \inf \left\{ k > 0 \mid \mathbb{E} \left[ \Phi \left( \frac{X}{k} \right) \right] \leq 1 \right\}.$$

One easily verifies that Orlicz premia are monotone, positively homogeneous, subadditive, law invariant, have the Lebesgue property, and satisfy  $H^\Phi(c) = c$ , for every  $c \geq 0$ . For  $\Phi(x) = x^p$ , with  $p \geq 1$ , clearly  $H^\Phi(X) = \|X\|_p$ . Orlicz premia are law-invariant norms and their natural domain is the nonnegative cone of an Orlicz space

$$L_+^\Phi := \left\{ X \geq 0 \mid \mathbb{E} \left[ \Phi \left( \frac{X}{k} \right) \right] < +\infty, \text{ for some } k > 0 \right\}.$$

We refer to Rao and Ren [37], Haezendonck and Goovaerts [31], Bellini and Rosazza Gianin [2, 3] and Cheridito and Li [11, 12] for further properties of Orlicz premia and Orlicz spaces. Notice that if  $X \in L_+^\infty$ ,  $X \neq 0$ , then

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H^\Phi(X)} \right) \right] = 1,$$

and moreover  $H^\Phi(X) = 1 \iff \mathbb{E}[\Phi(X)] = 1$ .

### 3 Dynamic Orlicz premia

In this section we extend the definition of Orlicz premia to a dynamic setting. On the probability space  $(\Omega, \mathcal{F}, P)$  we fix a filtration  $(\mathcal{F}_t)_{t \in [0, T]}$ , with  $\mathcal{F}_0 = \{\emptyset, \Omega\}$  and  $\mathcal{F}_T = \mathcal{F}$ .

**Definition 2.** Let  $\Phi$  be a Young function and let  $t, u \in [0, T]$  with  $t \leq u$ . For  $X \in L_+^\infty(\mathcal{F}_u)$ , the dynamic Orlicz premium  $H_t^\Phi: L_+^\infty(\mathcal{F}_u) \rightarrow L_+^\infty(\mathcal{F}_t)$  is

$$H_t^\Phi(X) := \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \mathbb{E}_P \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\}. \quad (3.1)$$

Clearly  $H_0^\Phi = H^\Phi$ . Recall that the essential infimum of a family of  $\mathcal{F}_t$ -measurable functions  $\{h_\alpha\}_{\alpha \in I}$  is the  $P$ -a.s. unique  $\mathcal{F}_t$ -measurable function  $Z$  such that  $Z \geq h_\alpha$  for each  $\alpha \in I$ , and if  $Z'$  is another  $\mathcal{F}_t$ -measurable function satisfying  $Z' \geq h_\alpha$  for each  $\alpha \in I$  then  $Z' \geq Z$  (see e.g., Föllmer and Schied [25]). Since  $X \in L_+^\infty(\mathcal{F}_u)$ , the set

$$\left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \mathbb{E}_P \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\}$$

is non void, so the definition is well-posed. The properties of  $H_t^\Phi$  are similar to those of  $H^\Phi$  and are reported in the following proposition.

**Proposition 3.** Let  $\Phi$  be a Young function, let  $s, t \in [0, T]$  with  $s \leq t$  and let  $H_t^\Phi: L_+^\infty(\mathcal{F}_T) \rightarrow L_+^\infty(\mathcal{F}_t)$  be defined as in (3.1). Then for each  $X, Y \in L_+^\infty(\mathcal{F}_T)$  it holds that:

- (a)  $X \leq Y \Rightarrow H_t^\Phi(X) \leq H_t^\Phi(Y)$
- (b)  $H_t^\Phi(X + Y) \leq H_t^\Phi(X) + H_t^\Phi(Y)$
- (c)  $X \in L_+^\infty(\mathcal{F}_t) \Rightarrow H_t^\Phi(X) = X$
- (d)  $H_t^\Phi(\lambda_s X) = \lambda_s H_t^\Phi(X), \forall \lambda_s \in L_{++}^\infty(\mathcal{F}_s)$
- (e)  $H_t^\Phi(X + \eta_s) \leq H_t^\Phi(X) + \eta_s, \forall \eta_s \in L_+^\infty(\mathcal{F}_s)$
- (f)  $\mathbb{E}_P[X | \mathcal{F}_t] \leq H_t^\Phi(X) \leq \|X\|_\infty$
- (g)  $A \in \mathcal{F}_t \Rightarrow H_t^\Phi(X 1_A) = 1_A H_t^\Phi(X)$
- (h) if  $X > 0$ , then

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H_t^\Phi(X)} \right) \mid \mathcal{F}_t \right] = 1$$

- (i)  $H_t^\Phi(X) = 1 \iff \mathbb{E}[\Phi(X) | \mathcal{F}_t] = 1$
- (j) if  $X_n \downarrow X$ , or if  $X_n \uparrow X$ , or if  $X_n \rightarrow X$  with  $\|X_n\| \leq k$ , then

$$H_t^\Phi(X_n) \rightarrow H_t^\Phi(X).$$

(k) if  $F_t$  is a regular version of the conditional distribution of  $X$  given  $\mathcal{F}_t$ , then

$$H_t^\Phi(X) = H^\Phi(F_t(\cdot, \omega))$$

*Proof.* The proof of (a) and (b) is straightforward and similar to the static case.

(c) If  $X \in L_{++}^\infty(\mathcal{F}_t)$  then from the properties of  $\Phi$  it follows that

$$\begin{aligned} & \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \mathbb{E}_P \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\} = \\ & \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \Phi \left( \frac{X}{h_t} \right) \leq 1 \right\} = \\ & \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid X \leq h_t \right\} = X \end{aligned}$$

(d) It holds that

$$\begin{aligned} & \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \mathbb{E}_P \left[ \Phi \left( \frac{\lambda_s X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\} = \\ & \text{ess inf} \left\{ \lambda_s h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \mathbb{E}_P \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\} = \\ & \lambda_s H_t^\Phi(X). \end{aligned}$$

(e) follows from (b) and (c).

(f) From the conditional Jensen inequality it follows that

$$\mathbb{E} \left[ \Phi \left( \frac{X}{\mathbb{E}[X|\mathcal{F}_t]} \right) \mid \mathcal{F}_t \right] \geq \Phi(1) = 1,$$

that gives the first part of the thesis. The second follows immediately from (a).

(g) follows immediately from (d).

(h) The set

$$A_t := \left\{ h_t \in L_{++}^\infty : E_P \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\}$$

is downward directed, hence there exists  $z_n \in A_t$ ,  $z_n \downarrow H_t^\Phi(X)$ ,  $P$ -a.s. From the dominated convergence theorem, it follows that

$$E_P \left[ \Phi \left( \frac{X}{H_t^\Phi(X)} \right) \mid \mathcal{F}_t \right] \leq 1.$$

Let now  $\bar{z}_n \in L_{++}^\infty$ , with  $\bar{z}_n < H_t^\Phi(X)$  and  $\bar{z}_n \uparrow H_t^\Phi(X)$ . Since  $H_t^\Phi(X) = \text{ess inf } A_t$ , it holds that

$$E_P \left[ \Phi \left( \frac{X}{\bar{z}_n} \right) \mid \mathcal{F}_t \right] > 1,$$

and again from the dominated convergence theorem,

$$E_P \left[ \Phi \left( \frac{X}{H_t^\Phi(X)} \right) \mid \mathcal{F}_t \right] \geq 1,$$

that gives the thesis.

(i) follows immediately from (h).

(j) Let  $X_n \downarrow X$ . Then by monotonicity  $H_n := H_t^\Phi(X_n) \downarrow H \geq H_t^\Phi(X)$ . It follows that

$$\mathbb{E}_P \left[ \Phi \left( \frac{X_n}{H} \right) \middle| \mathcal{F}_t \right] \geq 1,$$

and by the dominated convergence theorem

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H} \right) \middle| \mathcal{F}_t \right] \geq 1.$$

Similarly,

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H_n} \right) \middle| \mathcal{F}_t \right] \leq 1,$$

and by the dominated convergence theorem

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H} \right) \middle| \mathcal{F}_t \right] \leq 1,$$

so

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{H} \right) \middle| \mathcal{F}_t \right] = 1,$$

which from (h) implies that  $H = H_t^\Phi(X)$ . The proof of continuity from below is similar. To prove the last part of the thesis, let

$$Z_n = \sup_{k \geq n} X_k, Y_n = \inf_{k \geq n} X_k.$$

Then  $Z_n \geq X_n$ ,  $Z_n \downarrow X$  and  $Y_n \leq X_n$ ,  $Y_n \uparrow X$ , so from monotonicity and the first part of the thesis it follows that  $H_t^\Phi(X_n) \rightarrow H_t^\Phi(X)$ .

(k) Let  $F_t$  be a regular version of the conditional distribution of  $X$  given  $\mathcal{F}_t$ , that is let  $F_t: \mathbb{R} \times \Omega \rightarrow [0, 1]$  be such that for each  $\omega \in \Omega$ ,  $F_t(\cdot, \omega)$  is a distribution function on  $\mathbb{R}$  and for each  $x \in \mathbb{R}$  it holds  $F_t(x, \cdot) = \mathbb{E}_P[X \leq x | \mathcal{F}_t]$ . Since

$$\mathbb{E}_P \left[ \Phi \left( \frac{X}{h_t} \right) \middle| \mathcal{F}_t \right] = \int \Phi(x/h) dF_t(x, \omega) \text{ } P\text{-a.s.},$$

the thesis follows. □

**Example 4.** If  $\Phi(x) = x^p$  with  $p \geq 1$ , then

$$H_t^\Phi(X) = (\mathbb{E}[X^p | \mathcal{F}_t])^{1/p}.$$

## 4 Dynamic robust Orlicz premia

In Bellini et al. [4] we introduced robust Orlicz premia, arising from a penalized worst-case approach under ambiguity with respect to the true measure  $P$ . We considered two canonical cases of ambiguity averse preferences: variational preferences as in Maccheroni, Marinacci and Rustichini [36] and homothetic preferences as in Chateauneuf and Faro [8]. We recall the basic definitions and notations. We denote by  $\mathcal{Q}$  the set of all probability measures on  $(\Omega, \mathcal{F})$  that are absolutely continuous with respect to  $P$ .

**Definition 5.** Let  $\Phi$  be a Young function, let  $c: \mathcal{Q} \rightarrow [0, +\infty]$  be a penalty function satisfying  $\inf_{Q \in \mathcal{Q}} c(Q) = 0$ , and let  $\beta: \mathcal{Q} \rightarrow [0, 1]$  be a confidence function satisfying  $\sup_{Q \in \mathcal{Q}} \beta(Q) = 1$ . The robust Orlicz premia are defined by

$$H^{\Phi, c}(X) := \inf \left\{ k > 0 \mid \sup_{Q \in \mathcal{Q}} \mathbb{E}_Q \left[ \Phi \left( \frac{X}{k} \right) \right] - c(Q) \leq 1 \right\}.$$

$$H^{\Phi, \beta}(X) := \inf \left\{ k > 0 \mid \sup_{Q \in \mathcal{Q}} \mathbb{E}_Q \left[ \beta(Q) \Phi \left( \frac{X}{k} \right) \right] \leq 1 \right\}.$$

The corresponding dynamic robust Orlicz premia are defined as follows:

**Definition 6.** Let  $\Phi$  be a Young function, let  $\mathcal{C} = \{c_t\}_{t \in [0, T]}$  be a family of  $\mathcal{F}_t$ -measurable penalty functions  $c_t: \mathcal{Q} \rightarrow L_+^0(\mathcal{F}_t)$  satisfying  $\inf_{Q \in \mathcal{Q}} c_t(Q) = 0$ , and let  $\mathcal{B} = \{\beta_t\}_{t \in [0, T]}$  be a family of  $\mathcal{F}_t$ -measurable confidence functions  $\beta_t: \mathcal{Q} \rightarrow L_{[0, 1]}^0(\mathcal{F}_t)$  satisfying  $\sup_{Q \in \mathcal{Q}} \beta_t(Q) = 1$ , for each  $t \in [0, T]$ . Let  $t, u \in [0, T]$  with  $t \leq u$ . For  $X \in L_+^\infty(\mathcal{F}_u)$ , we define

$$H_t^{\Phi, \mathcal{C}}(X) := \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \text{ess sup}_{Q \in \mathcal{Q}} \left\{ \mathbb{E}_Q \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] - c_t(Q) \right\} \leq 1 \right\} \quad (4.2)$$

$$H_t^{\Phi, \mathcal{B}}(X) := \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \text{ess sup}_{Q \in \mathcal{Q}} \left\{ \mathbb{E}_Q \left[ \beta_t(Q) \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \right\} \leq 1 \right\} \quad (4.3)$$

Notice however that if, in (4.2), we define

$$\rho_t(X) := \text{ess sup}_{Q \in \mathcal{Q}} \{ \mathbb{E}_Q [X \mid \mathcal{F}_t] - c_t(Q) \}, \quad (4.4)$$

or similarly if, in (4.3), we define

$$\rho_t(X) := \text{ess sup}_{Q \in \mathcal{Q}} \{ \beta_t(Q) \mathbb{E}_Q [X \mid \mathcal{F}_t] \}, \quad (4.5)$$

then  $\rho_t$  is in both cases a dynamic risk measure that satisfies convexity, and (4.2) and (4.3) can be rewritten in a unified way as follows:

$$H_t^{\Phi, \rho}(X) := \text{ess inf} \left\{ h_t \in L_{++}^\infty(\mathcal{F}_t) \mid \rho_t \left( \Phi \left( \frac{X}{h_t} \right) \right) \leq 1 \right\}. \quad (4.6)$$

In other words, dynamic robust Orlicz premia arise by replacing the conditional expectation operator in (3.1) with a more general dynamic risk measure  $\rho_t$  that is convex. We begin by illustrating some simple special cases of Definition 6 in the case of variational preferences.

**Example 7.** Let  $c_t = 0$  for each  $t \in [0, T]$ . Then,  $\rho_t^{\mathcal{C}}(X) = \text{ess sup}[X | \mathcal{F}_t]$ , and

$$\begin{aligned} H_t^{\Phi, \mathcal{C}}(X) &= \text{ess inf} \left\{ h_t \in L_{++}^{\infty}(\mathcal{F}_t) \mid \text{ess sup} \left[ \Phi \left( \frac{X}{h_t} \right) \mid \mathcal{F}_t \right] \leq 1 \right\} \\ &= \text{ess sup}[X | \mathcal{F}_t]. \end{aligned}$$

**Example 8.** Let  $\Phi(x) = x^p$ , with  $p \geq 1$ , and let

$$c_t(Q) = \begin{cases} 0 & \text{if } Q \in \mathcal{S}_t \\ +\infty & \text{if } Q \notin \mathcal{S}_t, \end{cases}$$

with  $\mathcal{S}_t \subset \mathcal{Q}$ , for  $t \in [0, T]$ . Then,

$$H_t^{\Phi, \mathcal{C}}(X) = \text{ess sup}_{Q \in \mathcal{S}_t} (\mathbb{E}_Q [X^p | \mathcal{F}_t])^{1/p}.$$

In particular, for  $p = 1$  we get

$$H_t^{\Phi, \mathcal{C}}(X) = \text{ess sup}_{Q \in \mathcal{S}_t} \mathbb{E}_Q [X | \mathcal{F}_t],$$

that is a dynamic coherent risk measure in the usual sense.

**Example 9.** Let  $\Phi(x) = x^p$ , with a general  $\mathcal{C}$ . Then,

$$H_t^{\Phi, \mathcal{C}}(X) = \text{ess sup}_{Q \in \mathcal{Q}} \left( \frac{\mathbb{E}_Q [X^p | \mathcal{F}_t]}{1 + c_t(Q)} \right)^{1/p}.$$

Indeed, Definition 6 becomes

$$H_t^{\Phi, \mathcal{C}}(X) = \text{ess inf} \left\{ h_t \in L_{++}^{\infty}(\mathcal{F}_t) \mid \text{ess sup}_{Q \in \mathcal{Q}} \left\{ \frac{\mathbb{E}_Q [X^p | \mathcal{F}_t]}{h_t^p} - c_t(Q) \right\} \leq 1 \right\},$$

and the condition

$$\text{ess sup}_{Q \in \mathcal{Q}} \left\{ \frac{\mathbb{E}_Q [X^p | \mathcal{F}_t]}{h_t^p} - c_t(Q) \right\} \leq 1$$

is equivalent to

$$h_t \geq \text{ess sup}_{Q \in \mathcal{Q}} \left( \frac{\mathbb{E}_Q [X^p | \mathcal{F}_t]}{1 + c_t(Q)} \right)^{1/p}.$$

Notice that the quantity  $1/(1 + c_t(Q))$  is an  $\mathcal{F}_t$ -measurable random variable taking values in  $[0, 1]$ , so it may be interpreted as a discount factor.

Some of the properties of dynamic Orlicz premia remain valid also in the robust case; we list them in the following proposition.

**Proposition 10.** *Let  $\Phi$  be a Young function, let  $s, t \in [0, T]$  with  $s \leq t$ , and let  $H_t^{\Phi, \rho}: L_+^\infty(\mathcal{F}_T) \rightarrow L_+^\infty(\mathcal{F}_t)$  be as in (4.6). Then, for each  $X, Y \in L_+^\infty(\mathcal{F}_T)$ , it holds that:*

- (a)  $X \leq Y \Rightarrow H_t^{\Phi, \rho}(X) \leq H_t^{\Phi, \rho}(Y)$
- (b)  $H_t^{\Phi, \rho}(X + Y) \leq H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)$
- (c)  $X \in L_+^\infty(\mathcal{F}_t) \Rightarrow H_t^{\Phi, \rho}(X) = X$
- (d)  $H_t^{\Phi, \rho}(\lambda_s X) = \lambda_s H_t^{\Phi, \rho}(X), \forall \lambda_s \in L_+^\infty(\mathcal{F}_s)$
- (e)  $H_t^{\Phi, \rho}(X + \eta_s) \leq H_t^{\Phi, \rho}(X) + \eta_s, \forall \eta_s \in L_+^\infty(\mathcal{F}_s)$
- (f)  $H_t^{\Phi, \rho}(X) \leq \|X\|_\infty$
- (g)  $A \in \mathcal{F}_t \Rightarrow H_t^{\Phi, \rho}(X 1_A) = 1_A H_t^{\Phi, \rho}(X)$
- (h) if  $\rho_t$  has the Lebesgue property, then for  $X > 0$  it holds that

$$\rho_t \left( \Phi \left( \frac{X}{H_t^{\Phi, \rho}(X)} \right) \right) = 1$$

- (i) if  $\rho_t$  has the Lebesgue property, then  $H_t^{\Phi, \rho}(X) = 1 \iff \rho_t(\Phi(X)) = 1$
- (j) if  $\rho_t$  has the Lebesgue property, then if  $X_n \downarrow X$ , or if  $X_n \uparrow X$ , or if  $X_n \rightarrow X$  with  $\|X_n\| \leq k$ , it follows that

$$H_t^{\Phi, \rho}(X_n) \rightarrow H_t^{\Phi, \rho}(X)$$

- (k) if  $\rho_t$  is conditionally law-invariant, then also  $H_t^{\Phi, \rho}$  is conditionally law-invariant.

*Proof.* (a) follows immediately from the monotonicity of  $\rho_t$ . (b) Notice that

$$\begin{aligned} & \rho_t \left( \Phi \left( \frac{X + Y}{H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)} \right) \right) \\ & \leq \rho_t \left( \frac{H_t^{\Phi, \rho}(X)}{H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)} \Phi \left( \frac{X}{H_t^{\Phi, \rho}(X)} \right) + \frac{H_t^{\Phi, \rho}(Y)}{H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)} \Phi \left( \frac{Y}{H_t^{\Phi, \rho}(Y)} \right) \right) \\ & \leq \frac{H_t^{\Phi, \rho}(X)}{H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)} \rho_t \left( \Phi \left( \frac{X}{H_t^{\Phi, \rho}(X)} \right) \right) + \frac{H_t^{\Phi, \rho}(Y)}{H_t^{\Phi, \rho}(X) + H_t^{\Phi, \rho}(Y)} \rho_t \left( \Phi \left( \frac{Y}{H_t^{\Phi, \rho}(Y)} \right) \right) \\ & \leq 1, \end{aligned}$$

where the first inequality follows from the convexity of  $\Phi$  and the second from the convexity of  $\rho_t$ , from which the thesis follows.

(c) Notice that if  $X \in L_+^\infty(\mathcal{F}_t)$ , then  $\rho_t(X) = X$ , as a consequence of the hypothesis  $\inf_{Q \in \mathcal{Q}} c(Q) = 0$  and  $\sup_{Q \in \mathcal{Q}} \beta(Q) = 1$ . The thesis then follows as in item (c) of Proposition 3. The proofs of (d), (e), (f), (g), (h), (i), (j), (k) are identical to the corresponding items in Proposition 3, since under the assumption that  $\rho_t$  has the Lebesgue property, the dominated convergence theorem can still be applied.  $\square$

Summing up, the relevant properties of  $\rho_t$  in Proposition 10 are monotonicity (for (a)), convexity (for (b)), constancy (for (c)), the Lebesgue property (for (h), (i) and (j)), and conditional law-invariance (for (k)). In particular, neither translation invariance of  $\rho_t$  (which is satisfied in the case of variational preferences (4.4) but not in the case of homothetic preferences (4.5)), nor conditional positive homogeneity (which is satisfied in the case of homothetic preferences (4.5) but not in the case of variational preferences (4.4)) play a role in the proof of Proposition 10. Sufficient conditions on the penalty functions in (4.4) that guarantee the validity of these properties are well known in the literature on dynamic convex risk measures; we refer e.g., to Detlefsen and Scandolo [20], Föllmer and Schied [25]. The case of homothetic preferences is less explored.

## 5 Dynamic robust Haezendonck-Goovaerts risk measures

Dynamic robust Orlicz premia do not satisfy a translation invariance property, and are defined only for nonnegative losses. A construction via optimized translation-invariant extensions (Rockafellar and Uryasev [39] and Rockafellar, Uryasev and Zabarankin [40]) that resolves both issues has been suggested in Goovaerts et al. [30] (see also Bellini and Rosazza Gianin [2, 3]), leading to the so-called Haezendonck-Goovaerts risk measures (HG henceforth), of which a robust version has been introduced in Bellini et al. [4]. Their dynamic extension can be given as follows.

**Definition 11.** *Let  $\Phi$  be a Young function, let  $s, t \in [0, T]$  with  $s \leq t$ , and let  $H_t^{\Phi, \rho}$  be as in (4.6). For  $X \in L^\infty(\mathcal{F}_T)$ , we define*

$$HG_t^{\Phi, \rho}(X) := \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{x_t + H_t^{\Phi, \rho}((X - x_t)^+)\}. \quad (5.7)$$

All the properties of  $H_t^{\Phi, \rho}$  are inherited also by  $HG_t^{\Phi, \rho}$ .

**Proposition 12.** *Let  $\Phi$  be a Young function, let  $s, t \in [0, T]$  with  $s \leq t$ , and let  $H_t^{\Phi, \rho}$  be as in (4.6). Then, for each  $X, Y \in L^\infty(\mathcal{F}_T)$ , the following hold:*

- (a)  $X \leq Y \Rightarrow HG_t^{\Phi, \rho}(X) \leq HG_t^{\Phi, \rho}(Y)$
- (b)  $HG_t^{\Phi, \rho}(X + Y) \leq HG_t^{\Phi, \rho}(X) + HG_t^{\Phi, \rho}(Y)$
- (c)  $X \in L_+^\infty(\mathcal{F}_t) \Rightarrow HG_t^{\Phi, \rho}(X) = X$

- (d)  $HG_t^{\Phi, \rho}(\lambda_s X) = \lambda_s HG_t^{\Phi, \rho}(X)$ ,  $\forall \lambda_s \in L_{++}^\infty(\mathcal{F}_s)$   
(e)  $HG_t^{\Phi, \rho}(X + \eta_t) = HG_t^{\Phi, \rho}(X) + \eta_t$ ,  $\forall \eta_t \in L^\infty(\mathcal{F}_t)$   
(f)  $HG_t^{\Phi, \rho}(X) \leq \|X\|_\infty$   
(g)  $A \in \mathcal{F}_t \Rightarrow HG_t^{\Phi, \rho}(X 1_A) = 1_A HG_t^{\Phi, \rho}(X)$   
(h) if  $\rho_t$  satisfies the Lebesgue property, then

$$X_n \downarrow X \Rightarrow HG_t^{\Phi, \mathcal{C}}(X_n) \rightarrow HG_t^{\Phi, \mathcal{C}}(X), P\text{-a.s.}$$

- (i) if  $\rho_t$  is conditionally law-invariant, then also  $HG_t^{\Phi, \rho}$  is conditionally law-invariant.

*Proof.* (a) follows immediately from the monotonicity of  $H_t^{\Phi, \rho}$ . (b) follows from the subadditivity of the positive part and of  $H_t^{\Phi, \rho}$ . (c) If  $X \in L_+^\infty(\mathcal{F}_t)$ , then

$$HG_t^{\Phi, \rho}(X) = \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{x_t + (X - x_t)^+\} = X.$$

(d) If  $\lambda_s \in L_{++}^\infty(\mathcal{F}_s)$ , then

$$\begin{aligned} HG_t^{\Phi, \rho}(\lambda_s X) &= \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{x_t + H_t^{\Phi, \rho}((\lambda_s X - x_t)^+)\} = \\ &= \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{\lambda_s x_t + H_t^{\Phi, \rho}((\lambda_s X - \lambda_s x_t)^+)\} = \lambda_s HG_t^{\Phi, \rho}(X), \end{aligned}$$

from the conditional positive homogeneity of  $H_t^{\Phi, \rho}$ . (e) follows immediately from (5.7). (f) and (g) can be proved as in Proposition 10. (h) from Proposition 10, item (j), it follows that  $H_t^{\Phi, \rho}$  is continuous from above, so, if  $X_n \downarrow X$ , then

$$\begin{aligned} \inf_n HG_t^{\Phi, \rho}(X_n) &= \inf_n \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{x_t + H_t^{\Phi, \rho}((X_n - x_t)^+)\} \\ &= \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \inf_n \{x_t + H_t^{\Phi, \rho}((X_n - x_t)^+)\} \\ &= \operatorname{ess\,inf}_{x_t \in L^\infty(\mathcal{F}_t)} \{x_t + H_t^{\Phi, \rho}((X - x_t)^+)\} = HG_t^{\Phi, \rho}(X), \end{aligned}$$

from which the thesis follows. (i) can be proved as in the static case.  $\square$

The preceding proposition shows that dynamic robust HG risk measures are dynamic coherent risk measures in the sense of Delbaen [17] and Riedel [38], so they possess a dual representation in terms of essential suprema of conditional expectations, which is given in the following proposition. We denote by  $\mathcal{Q}_t \subset \mathcal{Q}$  the set of probability measures on  $(\Omega, \mathcal{F})$  such that  $Q = P$  on  $\mathcal{F}_t$ .

**Proposition 13.** *Let  $\Phi$  be a Young function, let  $t, u \in [0, T]$  with  $t \leq u$ , and let  $HG_t^{\Phi, \rho}$  be as in Definition 11, with  $\rho_t$  being continuous from above. Then, for each  $X \in L^\infty(\mathcal{F}_u)$ ,*

$$HG_t^{\Phi, \rho}(X) = \operatorname{ess\,sup}_{Q \in \mathcal{R}_t} \mathbb{E}_Q[X | \mathcal{F}_t], \quad (5.8)$$

where

$$\mathcal{R}_t := \{Q \in \mathcal{Q}_t \mid \mathbb{E}_Q [Z \mid \mathcal{F}_t] \leq H_t^{\Phi, \rho}(Z^+), \text{ for each } Z \in L^\infty(\mathcal{F}_t)\}.$$

*Proof.* By Detlefsen and Scandolo [20] (see also Delbaen [17] and Klöppel and Schweizer [32]) it follows that

$$HG_t^{\Phi, \rho}(X) = \operatorname{ess\,sup}_{Q \in \mathcal{R}_t} \mathbb{E}_Q [X \mid \mathcal{F}_t],$$

where  $\mathcal{R}_t = \{Q \in \mathcal{Q}_t \mid \mathbb{E}_Q [Y \mid \mathcal{F}_t] \leq HG_t^{\Phi, \rho}(Y), \text{ for each } Y \in L^\infty(\mathcal{F}_u)\}$ . Furthermore, from (5.7) it follows that

$$\mathbb{E}_Q [Y \mid \mathcal{F}_t] \leq HG_t^{\Phi, \rho}(Y), \quad \forall Y \in L^\infty(\mathcal{F}_u),$$

is equivalent to

$$\mathbb{E}_Q [Y - x_t \mid \mathcal{F}_t] \leq H_t^{\Phi, \rho}((Y - x_t)^+), \quad \forall x_t \in L^\infty(\mathcal{F}_t), Y \in L^\infty(\mathcal{F}_u),$$

and, moreover, to

$$\mathbb{E}_Q [Z \mid \mathcal{F}_t] \leq H_t^{\Phi, \rho}[Z^+], \quad \forall Z \in L^\infty(\mathcal{F}_u),$$

from which the thesis follows.  $\square$

## 6 Time-consistency properties

We start by recalling several definitions related to time-consistency that have been considered in the literature.

**Definition 14.** For  $t \in [0, T]$ , let  $\pi_t: L^\infty(\mathcal{F}_T) \rightarrow L^\infty(\mathcal{F}_t)$  be a dynamic risk measure. Let  $0 \leq s < t \leq T$ . We consider the following properties:

- i)  $\pi_t(X) \leq \pi_t(Y) \Rightarrow \pi_s(X) \leq \pi_s(Y)$
- ii)  $\pi_t(X) = \pi_t(Y) \Rightarrow \pi_s(X) = \pi_s(Y)$
- iii)  $\pi_s(\pi_t(X)) = \pi_s(X)$
- iv)  $\pi_t(X) \leq 0 \Rightarrow \pi_s(X) \leq 0$  or  $\pi_t(X) \geq 0 \Rightarrow \pi_s(X) \geq 0$

Property i) is the standard definition of time-consistency, while property iii) is commonly referred to as recursiveness. Properties i), ii) and iii) are known to be equivalent for dynamic monetary risk measures (see e.g., Föllmer and Schied [25]). As we show in the next lemma, they are also equivalent for monotone, conditionally positively homogeneous, normalized dynamic risk measures. We will therefore refer to i), ii) and iii) simply as time-consistency. Properties iv) are called weak acceptance consistency and weak rejection consistency, respectively.

**Lemma 15.** *Let  $\pi_t: L_{++}^\infty(\mathcal{F}_T) \rightarrow L_+^\infty(\mathcal{F}_t)$  be a monotone and conditionally positively homogeneous dynamic risk measure with  $\pi_t(1) = 1$ . Then properties i), ii), and iii) are equivalent.*

*Proof.* Clearly, i) implies ii). From conditional positive homogeneity and the normalization  $\pi_t(1) = 1$ , it follows that  $\pi_t(\pi_t(X)) = \pi_t(1 \cdot \pi_t(X)) = \pi_t(X)$ , hence ii) implies iii). Let now  $\pi_t(X) \leq \pi_t(Y)$ . From the monotonicity of  $\pi_s$  it follows that  $\pi_s(\pi_t(X)) \leq \pi_s(\pi_t(Y))$ , and from iii) it follows that  $\pi_s(X) \leq \pi_s(Y)$ , which gives i).  $\square$

It is straightforward to see that if  $\Phi(x) = x^p$  in Definition 2, then  $H_t^\Phi$  is time-consistent. Indeed, in this case (see Example 4)

$$H_t^\Phi(X) = (\mathbb{E}[X^p | \mathcal{F}_t])^{1/p},$$

so

$$H_s^\Phi(H_t^\Phi(X)) = (\mathbb{E}[(\mathbb{E}[X^p | \mathcal{F}_t]) | \mathcal{F}_s])^{1/p} = H_s^\Phi(X).$$

The next proposition shows that the converse also holds. Hence, the only dynamic Orlicz premia that satisfy time-consistency are the conditional certainty equivalents with  $\Phi$  a power function (i.e., conditional  $p$ -norms).

**Theorem 16.** *Let  $H_t^\Phi$  be as in Definition 2. If  $H_t^\Phi$  is time-consistent in the sense of Definition 14, then*

$$H_t^\Phi(X) = \Phi^{-1}(\mathbb{E}[\Phi(X) | \mathcal{F}_t]), \quad (6.9)$$

with  $\Phi(x) = x^p$ , for some  $p \geq 1$ .

*Proof.* We first prove that if  $H_t^\Phi$  is time-consistent, then

$$H_0^\Phi(X) = \Phi^{-1}(\mathbb{E}[\Phi(X)]).$$

Let  $X \in L_+^\infty(\mathcal{F}_t)$ . If  $H_0^\Phi(X) = 1$ , then by Proposition 3 item (i)

$$H_0^\Phi(X) = \Phi^{-1}(\mathbb{E}[\Phi(X)]).$$

If  $H_0^\Phi(X) \neq 1$ , then from the continuity of  $\Phi$  there exists  $z > 0$  such that

$$\frac{1}{2}\Phi(z) + \frac{1}{2}\Phi(H_0^\Phi(X)) = 1, \quad (6.10)$$

which implies that  $H_0^\Phi(Z) = 1$ , where  $Z$  is defined as follows:

$$Z := \begin{cases} z & \text{with prob. } 1/2, \\ H_0^\Phi(X) & \text{with prob. } 1/2. \end{cases}$$

Considering now the random variable

$$Y := \begin{cases} z & \text{with prob. } 1/2, \\ X & \text{with prob. } 1/2, \end{cases}$$

the property of time-consistency implies that  $H_0^\Phi(Y) = H_0^\Phi(Z) = 1$ , which gives

$$\frac{1}{2}\Phi(z) + \frac{1}{2}\mathbb{E}[\Phi(X)] = 1. \quad (6.11)$$

Upon comparing (6.11) with (6.10) we get that

$$\mathbb{E}[\Phi(X)] = \Phi(H_0^\Phi(X)).$$

The thesis for a general  $t$  follows from conditional law invariance.  $\square$

**Remark 17.** *To prove the above result we could also apply Theorem 1.4 of Kupper and Schachermayer [33]. Indeed,  $H_t^\Phi$  satisfies constancy (Proposition 3, item (c)), locality (Proposition 3, item (g)) and has the Fatou property (Proposition 3, item (k)), while  $\|\cdot\|_\infty$ -continuity of  $H_0^\Phi$  follows immediately from Proposition 3, items (a) and (e). Strict monotonicity of  $H_0^\Phi$  follows from Lemma 3 in Bellini and Rosazza Gianin [2]. All the hypotheses of Theorem 1.4 of Kupper and Schachermayer [33] are then satisfied by  $H_0^\Phi$  restricted on  $L_{++}^\infty$ . By applying Theorem 1.4 of Kupper and Schachermayer [33], it follows that  $H_t^\Phi$  is of the form*

$$H_t^\Phi = \ell^{-1} \mathbb{E}[\ell(X) \mid \mathcal{F}_t],$$

for some strictly increasing and continuous  $\ell$  and  $X \in L^\infty(0; +\infty)$ . Moreover,  $\ell$  reduces to a power function or a logarithmic function because of positive homogeneity of  $H_t^\Phi$  and by the De Finetti-Nagumo-Kolmogorov Theorem on the characterization of the expected utility functional (see De Finetti [15], and also Frittelli [26] and Laeven and Stadje [34]). Subadditivity of  $H_t^\Phi$  excludes the logarithmic case (see, for instance, Bellini et al. [4]). The case of  $H_0^\Phi$  on the whole  $L_+^\infty(\mathcal{F}_T)$  can be obtained by continuity. Assume indeed that  $X$  takes value 0 somewhere (say, in  $A \in \mathcal{F}_T$ ). Then  $X_n = \frac{1}{n}1_A + X1_{A^c} \rightarrow X$  in the  $\|\cdot\|_\infty$  norm, so by continuity it follows that  $H_0^\Phi(X_n) \rightarrow H_0^\Phi(X)$ . Once  $\ell$  is extended with continuity at 0, for  $t = 0$  (6.9) is true for any  $X \in L_+^\infty(\mathcal{F}_T)$ , while the case of a general  $t$  follows from time-consistency.

Let us now consider the case of dynamic robust Orlicz premia, defined by (4.6). First notice that if  $\Phi(x) = x^p$  for  $p \geq 1$  and if  $\rho_t$  is time-consistent, then the corresponding dynamic robust Orlicz premium  $H_t^{\Phi, \rho}(X) = (\rho_t(X^p))^{1/p}$  is time-consistent, since

$$\begin{aligned} H_s^{\Phi, \rho} \left( H_t^{\Phi, \rho}(X) \right) &= H_s^{\Phi, \rho} \left( (\rho_t(X^p))^{1/p} \right) = (\rho_s(\rho_t(X^p)))^{1/p} = \\ &= (\rho_s(X^p))^{1/p} = H_s^{\Phi, \rho}(X). \end{aligned}$$

The converse follows straightforwardly and is summarized in the following lemma:

**Lemma 18.** *Let  $H_t^{\Phi, \rho}$  be as in (4.6) with  $\Phi(x) = x^p$  for  $p \geq 1$ . If  $H_t^{\Phi, \rho}$  is time-consistent in the sense of Definition 14, then  $\rho_t$  is time-consistent.*

Time-consistency of  $\rho_t$  of the form (4.4) is well studied in the literature; see e.g., Bion-Nadal [5, 6] for a characterization via an additive cocycle property of the penalty function. Time-consistency of  $\rho_t$  of the form (4.5) is less explored.

The following results provide some connections (or ‘inheritance relations’) between time-consistency properties of dynamic robust Orlicz premia  $H_t^{\Phi,\rho}$  and the corresponding dynamic robust Haezendonck-Goovaerts risk measures  $HG_t^{\Phi,\rho}$ . In order to prove time-consistency of  $HG_t^{\Phi,\rho}$ , by Bion-Nadal [5, 6] and Delbaen [17], it suffices to verify  $m$ -stability of the set of generalized scenarios in its dual representation.

**Proposition 19.** *Let  $H_t^{\Phi,\rho}$  be a dynamic robust Orlicz premium as in (4.6) and let  $\rho_t$  satisfy the Lebesgue property. If, for each  $s, t \in [0, T]$  with  $s \leq t$  and for any  $X \in L^\infty(\mathcal{F}_T)$  it holds that*

$$H_s^{\Phi,\rho}(H_t^{\Phi,\rho}(X)) \leq H_s^{\Phi,\rho}(X), \quad (6.12)$$

then the corresponding dynamic robust risk measure  $HG_t^{\Phi,\rho}$  in (5.7) is time-consistent.

*Proof.* Denote by  $HG_{t,u}^{\Phi,\rho}$  the restriction of  $HG_t^{\Phi,\rho}$  to  $L^\infty(\mathcal{F}_u)$ . By Proposition 13,  $HG_{t,u}^{\Phi,\rho}$  has the following dual representation:

$$HG_{t,u}^{\Phi,\rho}(X) = \operatorname{ess\,sup}_{R \in \mathcal{R}_{t,u}} \mathbb{E}_R[X | \mathcal{F}_t], \quad X \in L^\infty(\mathcal{F}_u),$$

where

$$\mathcal{R}_{t,u} = \{R \in \mathcal{P}_{t,u} \mid \mathbb{E}_R[Z | \mathcal{F}_t] \leq H_{t,u}^{\Phi,\rho}(Z^+) \text{ for any } Z \in L^\infty(\mathcal{F}_u)\} \quad (6.13)$$

and  $\mathcal{P}_{t,u}$  is the subset of probability measures in  $\mathcal{P}_t$  that are defined on  $(\Omega, \mathcal{F}_u)$ . By Bion-Nadal [5, 6] and Delbaen [17], in order to prove time-consistency of  $HG_t^{\Phi,\rho}$  it is sufficient to verify the  $m$ -stability of the dual sets  $(\mathcal{R}_{t,u})_{0 \leq t \leq u \leq T}$ . Let  $0 \leq s \leq t \leq u \leq T$  and let  $R_1 \in \mathcal{R}_{s,t}$ ,  $R_2 \in \mathcal{R}_{t,u}$ . Denote by  $\bar{R}$  the pasting between  $R_1$  and  $R_2$ . To prove that  $\bar{R}$  belongs to  $\mathcal{R}_{s,u}$ , notice that from (6.12) it follows that for any  $Y \in L^\infty(\mathcal{F}_u)$

$$\begin{aligned} \mathbb{E}_{\bar{R}}[Y | \mathcal{F}_s] &= \mathbb{E}_{R_1}[\mathbb{E}_{R_2}[Y | \mathcal{F}_t] | \mathcal{F}_s] \leq \mathbb{E}_{R_1}[H_t^{\Phi,\rho}(Y^+) | \mathcal{F}_s] \\ &\leq H_s^{\Phi,\rho}(H_t^{\Phi,\rho}(Y^+)) \leq H_s^{\Phi,\rho}(Y^+), \end{aligned}$$

that implies that  $\bar{R} \in \mathcal{R}_{s,u}$ .  $\square$

In particular, the hypothesis (6.12) in Proposition 19 is satisfied if  $H_t^{\Phi,\rho}$  is time-consistent. A weaker version of the thesis holds in much more general situations, as the following proposition shows.

**Proposition 20.** *Let  $(H_t)_{t \in [0, T]}$  be any time-consistent family of functionals  $H_t: L_+^\infty(\mathcal{F}_T) \rightarrow L_+^\infty(\mathcal{F}_t)$  satisfying constancy and cash-subadditivity. Then the corresponding dynamic risk measure  $(HG_t)_{t \in [0, T]}$  defined as in (5.7) is weakly rejection consistent in the sense of Definition 14, that is, for any  $0 \leq s \leq t \leq T$ ,*

$$HG_t(X) \geq 0 \Rightarrow HG_s(X) \geq 0.$$

*Proof.* Suppose that  $HG_t(X) \geq 0$  for some  $X \in L^\infty(\mathcal{F}_T)$ . By (5.7) it follows that

$$x_t + H_t((X - x_t)^+) \geq 0, \quad \forall x_t \in L^\infty(\mathcal{F}_t),$$

hence, for any  $s \leq t$ ,

$$x_s + H_t((X - x_s)^+) \geq 0, \quad \forall x_s \in L^\infty(\mathcal{F}_s). \quad (6.14)$$

Fix now any  $s \leq t$ . We are going to prove that  $x_s + H_s((X - x_s)^+) \geq 0$  for any  $x_s \in L^\infty(\mathcal{F}_s)$ , hence  $HG_s \geq 0$ . By time-consistency and constancy of  $H_t$ , for any  $x_s \in L^\infty(\mathcal{F}_s)$  inequality (6.14) becomes

$$\begin{aligned} H_s(H_t((X - x_s)^+)) &\geq H_s(-x_s) \\ H_s((X - x_s)^+) &\geq -x_s \\ x_s + H_s((X - x_s)^+) &\geq 0. \end{aligned} \quad (6.15)$$

A fortiori, inequality (6.15) is satisfied also for any  $x_s \in L_+^\infty(\mathcal{F}_s)$ . For an arbitrary  $x_s \in L^\infty(\mathcal{F}_s)$ , set  $A = \{x_s \geq 0\} \in \mathcal{F}_s$ . Inequality (6.14) implies therefore that

$$x_s 1_A + H_t((X - x_s)^+) \geq -x_s 1_{A^c}.$$

Proceeding as above, we get

$$\begin{aligned} H_s(x_s 1_A + H_t((X - x_s)^+)) &\geq H_s(-x_s 1_{A^c}) \\ H_s(x_s 1_A) + H_s(H_t((X - x_s)^+)) &\geq -x_s 1_{A^c} \end{aligned} \quad (6.16)$$

$$x_s 1_A + H_s((X - x_s)^+) \geq -x_s 1_{A^c} \quad (6.17)$$

$$x_s + H_s((X - x_s)^+) \geq 0,$$

where (6.16) follows from subadditivity and constancy, while (6.17) follows from constancy and time-consistency of  $H_s$ . From the arguments above it follows that  $x_s + H_s((X - x_s)^+) \geq 0$  holds for any  $x_s \in L^\infty(\mathcal{F}_s)$ , hence  $HG_s(X) \geq 0$ .  $\square$

The following result is a partial converse of Proposition 20. We omit the proof, which follows immediately by (5.7).

**Proposition 21.** *Let  $(H_t)_{t \in [0, T]}$  be a family of functionals  $H_t : L_+^\infty(\mathcal{F}_T) \rightarrow L_+^\infty(\mathcal{F}_t)$  satisfying monotonicity and constancy. If the corresponding dynamic risk measure  $HG_t$  satisfies weak rejection consistency in the sense of Definition 14 item v), then if for each  $x_t \in L^\infty(\mathcal{F}_t)$  it holds that*

$$H_t((Y - x_t)^+) \geq H_t((-x_t)^+),$$

then for each  $x_s \in L^\infty(\mathcal{F}_s)$ ,  $0 \leq s \leq t$ , it holds that

$$H_s((Y - x_s)^+) \geq H_s((-x_s)^+).$$

Notice that the thesis of Proposition 21 is weaker than and is implied by time-consistency of  $H_t$ . Indeed,

$$H_t((Y - x_t)^+) \geq H_t((-x_t)^+) \text{ for any } x_t \in L^\infty(\mathcal{F}_t)$$

implies that

$$H_t((Y - x_s)^+) \geq H_t((-x_s)^+) \text{ for any } x_s \in L^\infty(\mathcal{F}_s),$$

and by monotonicity and time-consistency of  $H_s$  it follows that

$$H_s((Y - x_s)^+) = H_s(H_t((Y - x_s)^+)) \geq H_s(H_t((-x_s)^+)) = H_s((-x_s)^+),$$

for any  $x_s \in L^\infty(\mathcal{F}_s)$ .

## References

- [1] ARTZNER, P., F. DELBAEN, J.M. EBER, D. HEATH AND H. KU (2004). Coherent multiperiod risk adjusted values and Bellman's principle. Preprint, ETH Zurich, Switzerland.
- [2] BELLINI, F. AND E. ROSAZZA GIANIN (2008). On Haezendonck risk measures. *Journal of Banking and Finance* 32, 986-994.
- [3] BELLINI, F. AND E. ROSAZZA GIANIN (2012). Haezendonck-Goovaerts risk measures and Orlicz quantiles. *Insurance: Mathematics and Economics* 51, 107-114.
- [4] BELLINI, F., R.J.A. LAEVEN AND E. ROSAZZA GIANIN (2018). Robust return risk measures. *Mathematics and Financial Economics* 12, 5-32.
- [5] BION-NADAL, J. (2008). Dynamic risk measures: Time consistency and risk measures from BMO martingales. *Finance and Stochastics* 12, 219-244.
- [6] BION-NADAL, J. (2009). Time consistent dynamic risk processes. *Stochastic Processes and their Applications* 119, 633-654.
- [7] CERREIA-VIOGLIO, S., F. MACCHERONI, M. MARINACCI AND L. MONTRUCCHIO (2011). Uncertainty averse preferences. *Journal of Economic Theory* 146, 1275-1330.
- [8] CHATEAUNEUF, A. AND J.H. FARO (2010). Ambiguity through confidence functions. *Journal of Mathematical Economics* 45, 535-558.
- [9] CHEN, Z. AND L.G. EPSTEIN (2002). Ambiguity, risk, and asset returns in continuous time. *Econometrica* 70, 1403-1443.
- [10] CHERIDITO, P., F. DELBAEN AND M. KUPPER (2006). Dynamic monetary risk measures for bounded discrete time processes. *Electronic Journal of Probability* 11, 57-106.

- [11] CHERIDITO, P. AND T. LI (2008). Dual characterization of properties of risk measures on Orlicz hearts. *Mathematics and Financial Economics* 2, 29-55.
- [12] CHERIDITO, P. AND T. LI (2009). Risk measures on Orlicz hearts. *Mathematical Finance* 19, 189-214.
- [13] CHERIDITO, P. AND M. KUPPER (2009). Recursiveness of indifference prices and translation-invariant preferences. *Mathematics and Financial Economics* 2, 173-188.
- [14] CHERIDITO, P. AND M. KUPPER (2011). Composition of time-consistent dynamic monetary risk measures in discrete time. *International Journal of Theoretical and Applied Finance* 14, 137-162.
- [15] DE FINETTI, B. (1931). Sul concetto di media. *Giornale degli Economisti* 2, 369-396.
- [16] DELBAEN, F. (2002). Coherent risk measures on general probability spaces. In: Sandmann, K., Schönbucher, P.J. (Eds.), *Advances in Finance and Stochastics: Essays in Honour of Dieter Sondermann*, Springer, Berlin, pp. 1-37.
- [17] DELBAEN, F. (2006). The structure of  $m$ -stable sets and in particular of the set of risk neutral measures. In: *Memoriam Paul-André Meyer, Lecture Notes in Mathematics* 1874, pp. 215-258.
- [18] DELBAEN, F., S. PENG AND E. ROSAZZA GIANIN (2010). Representation of the penalty term of dynamic concave utilities. *Finance and Stochastics* 14, 449-472.
- [19] DELBAEN, F., F. BELLINI, V. BIGNOZZI AND J. ZIEGEL (2014). Risk measures with the CxLS property. <http://arxiv.org/abs/1411.0426>.
- [20] DETLEFSEN, K. AND G. SCANDOLO (2005). Conditional and dynamic convex risk measures. *Finance and Stochastics* 9, 539-561.
- [21] DUFFIE, D. AND L.G. EPSTEIN (1992). Stochastic differential utility. *Econometrica* 60, 353-394.
- [22] EPSTEIN, L.G. AND M. SCHNEIDER (2003). Recursive multiple-priors. *Journal of Economic Theory* 113, 1-31.
- [23] FÖLLMER, H. AND I. PENNER (2006). Convex risk measures and the dynamics of their penalty functions. *Statistics and Decisions* 14, 1-15.
- [24] FÖLLMER, H. AND A. SCHIED (2002). Convex measures of risk and trading constraints. *Finance & Stochastics* 6, 429-447.
- [25] FÖLLMER, H. AND A. SCHIED (2011). *Stochastic Finance*. 3rd ed., De Gruyter, Berlin.

- [26] FRITTELLI, M. (2000). Introduction to a theory of value coherent with the no-arbitrage principle. *Finance and Stochastics* 4, 275-297.
- [27] FRITTELLI, M. AND E. ROSAZZA GIANIN (2002). Putting order in risk measures. *Journal of Banking and Finance* 26, 1473-1486.
- [28] FRITTELLI, M. AND E. ROSAZZA GIANIN (2004). Dynamic convex risk measures. In: Szegö, G. (Ed.), *Risk Measures for the 21st Century*, J. Wiley, 227-248.
- [29] GILBOA, I. AND D. SCHMEIDLER (1989). Maxmin expected utility with non-unique prior. *Journal of Mathematical Economics* 18, 141-153.
- [30] GOOVAERTS, M.J., R. KAAS, J. DHAENE AND Q. TANG (2004). Some new classes of consistent risk measures. *Insurance: Mathematics and Economics* 34, 505-516.
- [31] HAEZENDONCK, J. AND M.J. GOOVAERTS (1982). A new premium calculation principle based on Orlicz norms. *Insurance: Mathematics and Economics* 1, 41-53.
- [32] KLÖPPEL, S. AND M. SCHWEIZER (2007). Dynamic utility indifference valuation via convex risk measures. *Mathematical Finance* 17, 599-627.
- [33] KUPPER, M. AND W. SCHACHERMAYER (2009). Representation results for law invariant time consistent functions. *Mathematics and Financial Economics* 2, 189-210.
- [34] LAEVEN, R.J.A. AND M.A. STADJE (2013). Entropy coherent and entropy convex measures of risk. *Mathematics of Operations Research* 38, 265-293.
- [35] LAEVEN, R.J.A. AND M.A. STADJE (2014). Robust portfolio choice and indifference valuation. *Mathematics of Operations Research* 39, 1109-1141.
- [36] MACCHERONI, F., M. MARINACCI AND A. RUSTICHINI (2006). Ambiguity aversion, robustness, and the variational representation of preferences. *Econometrica* 74, 1447-1498.
- [37] RAO, M.M. AND Z.D. REN (1991). *Theory of Orlicz Spaces*, Marcel Dekker, New York.
- [38] RIEDEL F. (2004). Dynamic coherent risk measures. *Stochastic Processes and their Applications* 112, 185-200.
- [39] ROCKAFELLAR, R.T. AND S.P. URYASEV (2000). Optimization of conditional value-at-risk. *Journal of Risk* 2, 21-42.
- [40] ROCKAFELLAR, R.T., S.P. URYASEV AND M. ZABARANKIN (2008). Risk tuning with generalized linear regression. *Mathematics of Operations Research* 33, 712-729.

- [41] RUSZCZYŃSKI, A. AND A. SHAPIRO (2006). Optimization of convex risk functions. *Mathematics of Operations Research* 31, 433-452.
- [42] RUSZCZYŃSKI, A. AND A. SHAPIRO (2006). Conditional risk mappings. *Mathematics of Operations Research* 31, 544-561.
- [43] TANG, S. AND W. WEI (2012). Representation of dynamic time-consistent convex risk measures with jumps. *Risk and Decision Analysis* 3, 167-190.
- [44] WEBER, S. (2006). Distribution-invariant risk measures, information, and dynamic consistency. *Mathematical Finance* 16, 419-442.